

STAT 5200 Handout #10a

Contrasts in Two-Way Factorial Design (Ch. 9)

Main effects with contrasts: (See HO # 9)

↳ interaction driven by Alcohol = 1

- Best to use effects parameterization to define contrasts here:

$$Y_{ijk} = \mu + A_i + B_j + AB_{ij} + \epsilon_{ijk} \quad k = 1 \dots 4$$

↑ Alcohol $i=1,2,3$ ↑ interaction $j=1,2$

- After accounting for effect of Base, does Alcohol have any effect?

Type III SS do this - test terms (like A) after accounting for (or while controlling for) all other terms in the model

- $H_0: A_1 = A_2 = A_3 = 0$

↳ by $\sum_i A_i = 0$ constraint

- This is actually the intersection of multiple nulls:

2 DF. $\left\{ \begin{array}{l} H_0^{(1)}: A_1 = A_2 \\ \text{and} \\ H_0^{(2)}: A_1 = A_3 \end{array} \right.$

-or- $A_1 = A_2$ and $A_2 = A_3$
-or- $A_2 = A_3$ and $A_1 = A_3$

- Think of these as multiple contrasts:

$H_0^{(1)}: \psi^{(1)} = 0$ where $\psi^{(1)} = A_1 - A_2$

and $H_0^{(2)}: \psi^{(2)} = 0$ where $\psi^{(2)} = A_1 - A_3$

(see pp. 5-6 of HO #9)

- We can similarly test effect of Base, after accounting for effect of Alcohol:

Contrasts to test interactions: (See HO # 10; T = Temp, S = Sucrose)

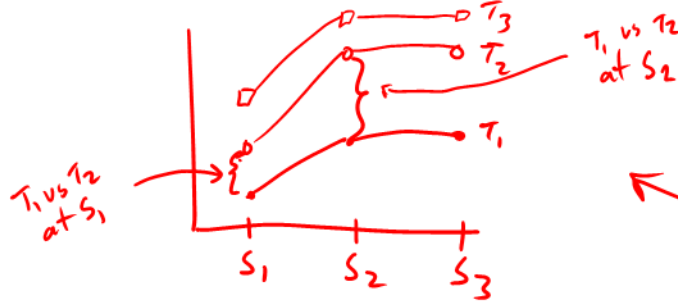
signif. T x S interaction is "driven" by Temp=20 / Sucr=20 combination

- Best to use effects parameterization to define contrasts here:

$$Y_{ijk} = \mu + T_i + S_j + TS_{ij} + \epsilon_{ijk}$$

$i=1..3$
 $j=1..3$
 $k=1..3$

- In testing interaction, look carefully at interaction plot:



- H_0 : there is no interaction involving T & S

some $\rightarrow H_0$: the effect of T does not depend on S (and vice-versa)

some $\rightarrow H_0$: the T-level comparisons are the same for each S level

T_1 vs T_2
 and
 T_1 vs T_3

S_1
 S_2
 S_3

no - profiles/lines aren't roughly parallel

- Consider this as the intersection of multiple hypotheses:

4 d.f. $\left\{ \begin{array}{l} H_0^{(1)} : TS_{11} - TS_{21} = TS_{12} - TS_{22} \quad (T_1 - T_2 \text{ is same at S levels } 1 \& 2) \\ H_0^{(2)} : TS_{11} - TS_{31} = TS_{12} - TS_{32} \quad (T_1 - T_3 \text{ is same at S levels } 1 \& 2) \\ H_0^{(3)} : TS_{11} - TS_{21} = TS_{13} - TS_{23} \quad (T_1 - T_2 \text{ is same at S levels } 1 \& 3) \\ H_0^{(4)} : TS_{11} - TS_{31} = TS_{13} - TS_{33} \quad (T_1 - T_3 \text{ is same at S levels } 1 \& 3) \end{array} \right.$

- What about the "vice-versa"?

- Note that $H_0^{(1)}$ can be rearranged as $TS_{11} - TS_{12} = TS_{21} - TS_{22}$
 ($S_1 - S_2$ is the same at T levels 1 and 2)

- Similarly:

- * $H_0^{(2)}$ is $S_1 - S_2$ at T levels 1 & 3
- * $H_0^{(3)}$ is $S_1 - S_3$ at T levels 1 & 2
- * $H_0^{(4)}$ is $S_1 - S_3$ at T levels 1 & 3



- So we have multiple contrasts:

$$\begin{aligned}\psi^{(1)} &= TS_{11} - TS_{21} - TS_{12} + TS_{22} \\ \psi^{(2)} &= TS_{11} - TS_{31} - TS_{12} + TS_{32} \\ \psi^{(3)} &= TS_{11} - TS_{21} - TS_{13} + TS_{23} \\ \psi^{(4)} &= TS_{11} - TS_{31} - TS_{13} + TS_{33}\end{aligned}$$

- Note that these contrasts are zero if and only if lines in interaction plot are parallel

(i.e., if! only if no interaction)

- So testing H_0 : "no interaction involving T & S" is equivalent to testing:

$$H_0^{(1)}: \psi^{(1)} = 0 \text{ and } H_0^{(2)}: \psi^{(2)} = 0 \text{ and } H_0^{(3)}: \psi^{(3)} = 0 \text{ and } H_0^{(4)}: \psi^{(4)} = 0$$

- Then can do these as post-hoc tests to characterize non-parallel nature of plot (or to check specific hypotheses) *(p.7 of H0 #10)*

- Contrasts involving interactions (or other custom comparisons) are notoriously awkward.

Without some thought and organization, CONTRAST statement in SAS will fail. *(no output!)*

Think of interaction subscripts being ordered as (row, column). Consider $\psi^{(1)}$ above:

"row" term is first in CLASS statement

	S		
	1	2	3
T	1	-1	0
2	-1	1	0
3	0	0	0

T level ↑ S level

← coeffs. in $\psi^{(1)}$, and in first block of p.7 of H0 #10; could also write as

$$1 \ -1 \ 0 \ -1 \ 1 \ 0 \ 0 \ 0 \ 0$$

(contrast not "estimable")

Contrasts for custom comparisons

- Example: In Handout #10 (T=Temp, S=Sucrose), compare Temp20/Sucrose60 with Temp30/Sucrose20

- General steps to build a meaningful CONTRAST statement:

1. Write out H_0 in terms of μ_{ij} 's (from means model)

$$H_0: \mu_{13} = \mu_{21}$$

2. Substitute in effects model (to "talk" to SAS)

$$\begin{aligned}\mu_{13} &= \mu + T_1 + S_3 + TS_{13} \\ \mu_{21} &= \mu + T_2 + S_1 + TS_{21}\end{aligned}$$

$$H_0: (\mu + T_1 + S_3 + TS_{13}) = (\mu + T_2 + S_1 + TS_{21})$$

3. Simplify to $H_0: \psi = 0$ (in terms of effect model parameters)

$$H_0: T_1 - T_2 - S_1 + S_3 + TS_{13} - TS_{21} = 0$$

ψ

even though only looks at means of 2 of 9 (3x3) factor level comb, the larger sample size & full structure of design will give greater power to detect a difference.

4. Use coefficients of ψ to define CONTRAST coefficients (in order)

$$T \quad 1 \quad -1 \quad 0 \quad S \quad -1 \quad 0 \quad 1$$

$$T \times S \quad \begin{bmatrix} 0 & 0 & 1 \\ -1 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}$$

row = i corresp. to T_i (row term is first in CLASS statement)
 col = j corresp. to S_j
 T_i S_1, S_2, S_3

- Custom comparison example: Does average of Temp20 (over all Sucrose levels) differ significantly from Temp30/Sucrose20? (See top plot on p. 3 of Handout #10)

μ_{ij} = mean at T level i & S level j ; $\mu_{ij} = \underbrace{\mu + T_i + S_j}_{\text{mems model}} + \underbrace{TS_{ij}}_{\text{effects model}}$

subs. in effects param.

$$H_0: \frac{\mu_{11} + \mu_{12} + \mu_{13}}{3} = \mu_{21}$$

$$H_0: (\mu + T_1 + S_1 + TS_{11}) + (\mu + T_1 + S_2 + TS_{12}) + (\mu + T_1 + S_3 + TS_{13}) = 3(\mu + T_2 + S_1 + TS_{21})$$

simplify

$$H_0: \underbrace{3T_1 - 3T_2 - 2S_1 + S_2 + S_3 + TS_{11} + TS_{12} + TS_{13} - 3TS_{21}}_{\psi} = 0$$

SAS: $T \quad 3 \quad -3 \quad 0 \quad S \quad -2 \quad 1 \quad 1 \quad T \times S \quad \begin{bmatrix} 1 & 1 & 1 \\ -3 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} ;$

- Custom comparison example: Does average of Temp40 over Sucrose20 and Sucrose40 differ significantly from average of Temp30 over Sucrose40 and Sucrose60? (See bottom plot on p. 3 of Handout #10)